

# Wishart spectral dynamics

Peter H. Gruber<sup>1</sup>, Claudio Tebaldi<sup>2</sup> and Fabio Trojani<sup>1</sup>

(1) Università della Svizzera Italiana, Lugano; (2) Università L. Bocconi, Milano

June 21, 2009

Workshop on Spectral and Cubature Methods in Finance and Econometrics, Leicester

# Statement of Purposes

- Introduction
  - Statement of purposes
- Plan of the talk
- Empirical Stylized Facts
- Model Definition and Identification
- Estimation
- Results
- Conclusion

Wishart pricing technology is appealing for many applications, Wishart efficient estimation is still under debate. The price of modeling flexibility is potentially paid in terms of estimation complexity and lack of identifiability.

- We focus on plain vanilla index option prices as provided in OptionMetrics database,
- Estimation shows a competitive pricing error using *one* set of parameters for *five years* of daily data and with diffusive processes
- Dynamic factors are identified in terms of observable portfolios: level, term structure slope and stochastic leverage.

# Plan of the talk

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## **Financial Motivation**

Preliminary PCA analysis of the IV surface

## **Model**

Framework, structure

## **Identification**

Spectral representation

## **Estimation**

Est. procedure, Computational Remarks

## **Results**

Pricing performance, a Stylized Model

## **Conclusion**

Conclusions and Perspectives

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PCA of IV surface

PCs

IV portfolios

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# Empirical Stylized Facts

# PCA of implied volatility surface

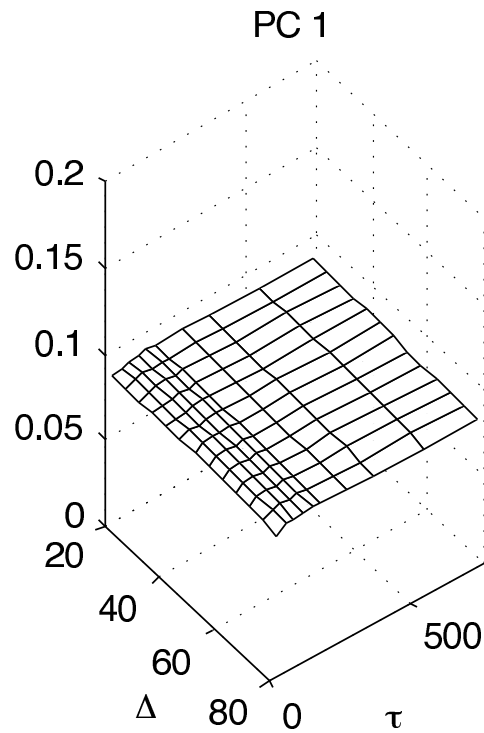
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- Data** S&P 500 call options
- 1996-2007/06 for explorative analysis
  - 2000/01-2004/12 for estimation
- Method** Plain PCA vectorized implied volatility matrix
- Grid of synthetic options from OptionMetrics
  - Delta range = 20%, 25%, . . . , 80%
  - maturities = 1,2,3,4,5,6,9,12,18,24 months

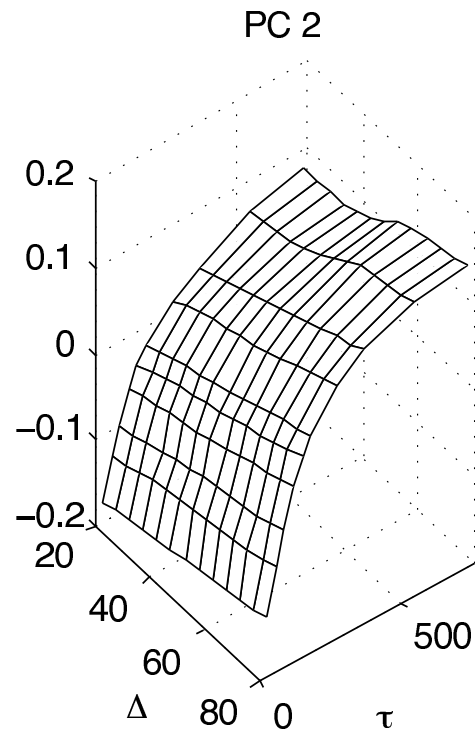
Sample	$l$	$PC_1$	$PC_2$	$PC_3$	$PC_4$	$n$
Full	2	95.46	3.32	0.67	0.17	2844
Low vola	3	93.24	4.65	0.98	0.46	1422
High vola	3	87.57	9.72	1.42	0.45	1422

No. of factors; Fraction explained

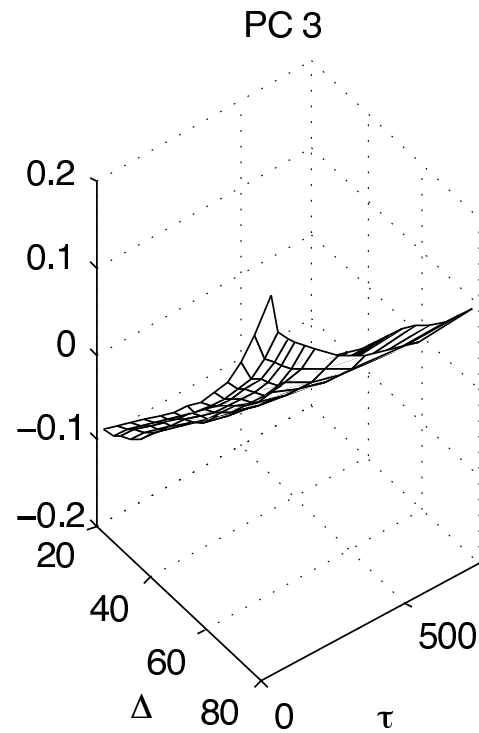
# Eigenvectors



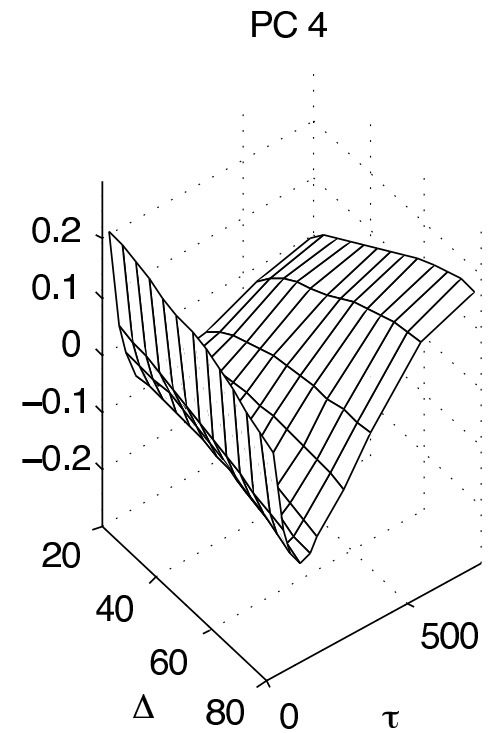
Level



Term structure



Risk reversal



TS hump

# Principal components

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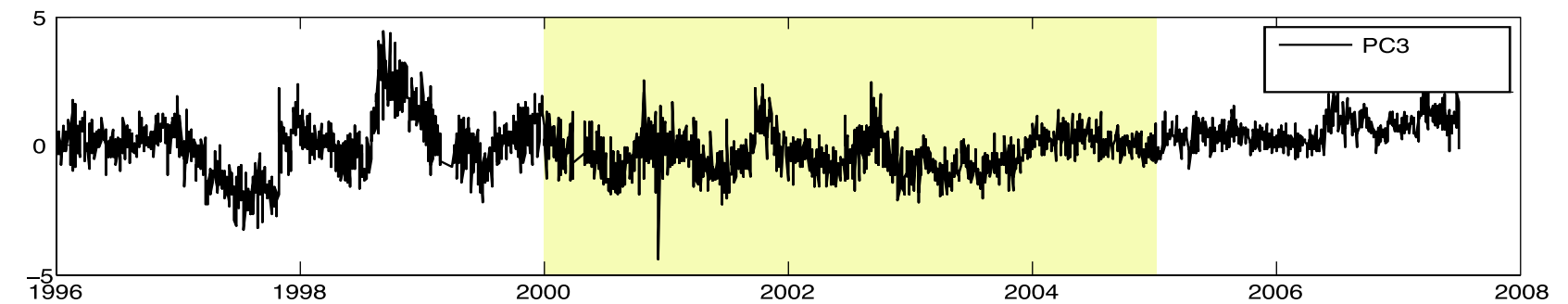
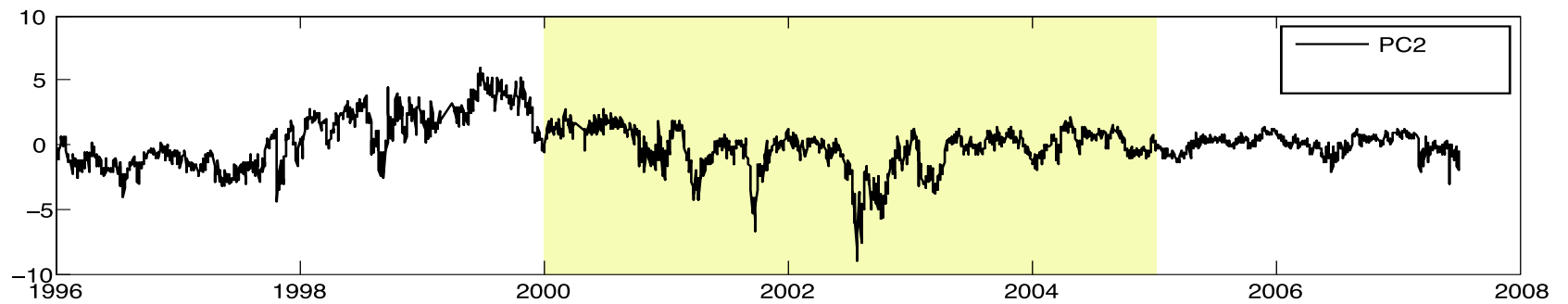
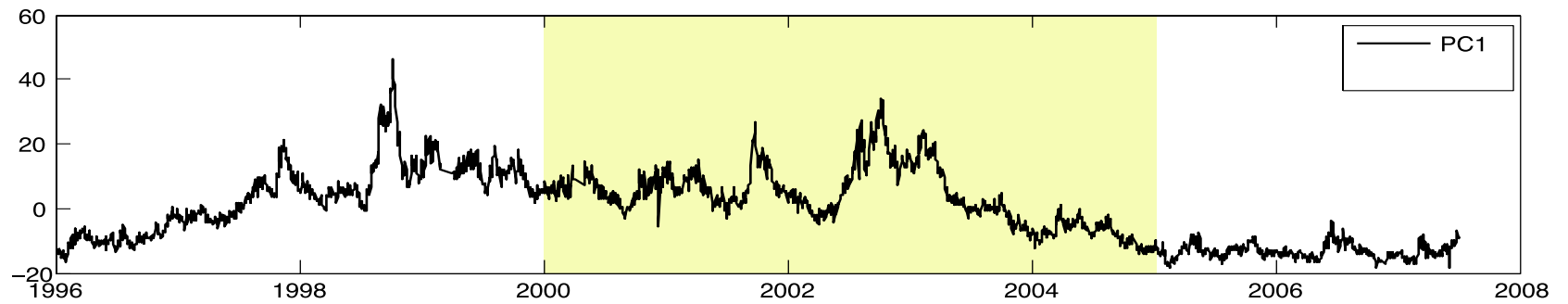
IV portfolios

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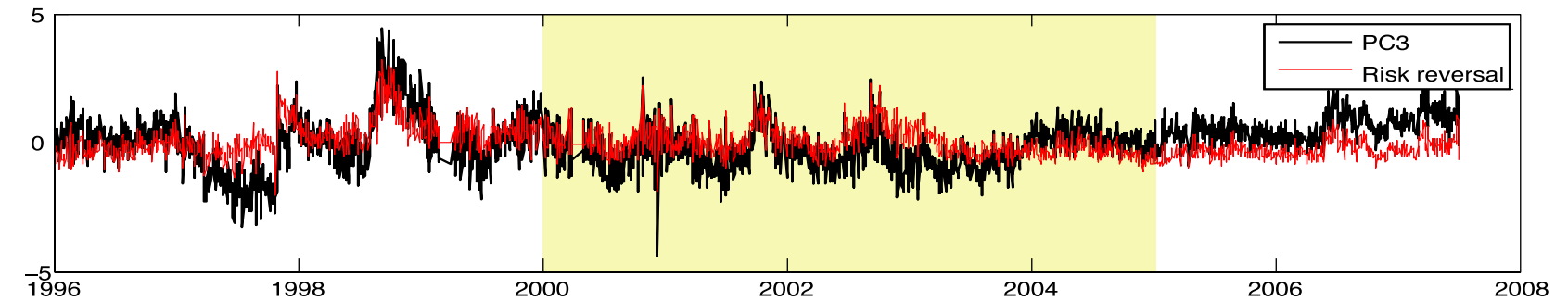
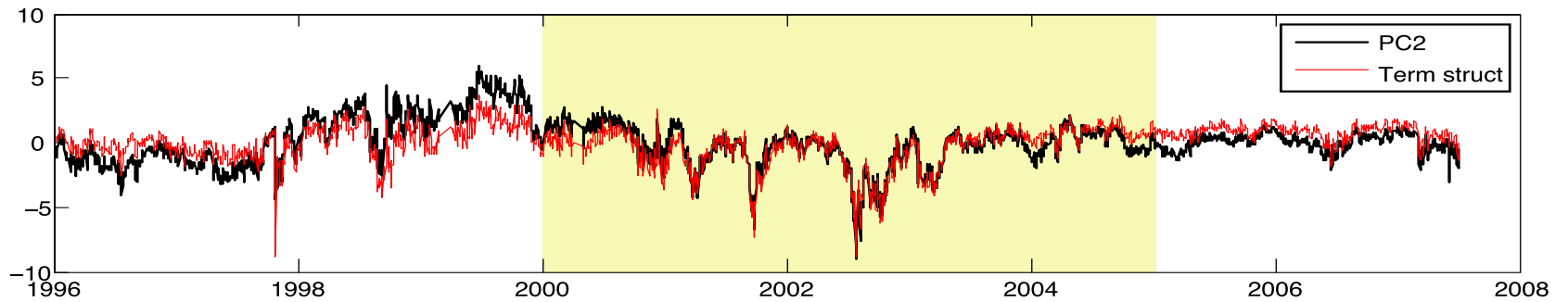
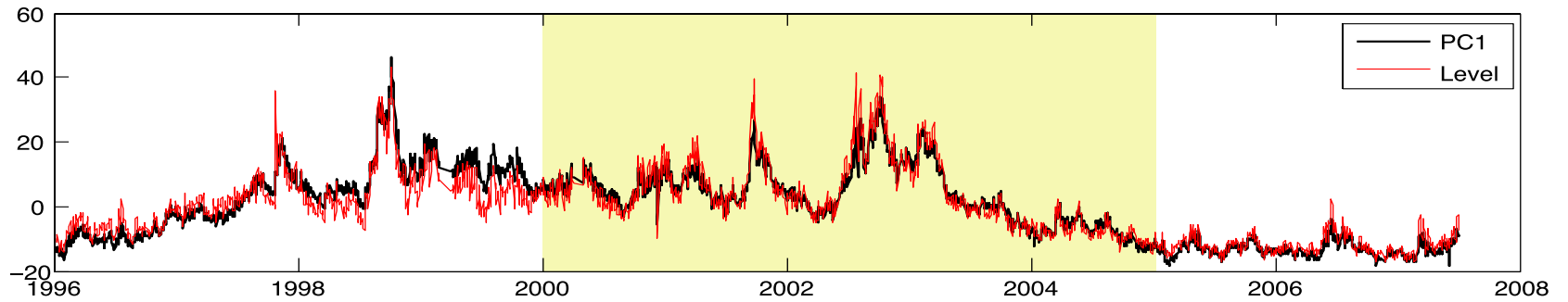


# IV portfolios

Introduction	(1) Level	$IV_{\Delta=50,\tau=30}$
Empirical Stylized Facts	(2) Term structure	$IV_{\Delta=50,\tau=30} - IV_{\Delta=50,\tau=365}$
PCA of IV surface	(3) RiskReversal	$IV_{\Delta=75,\tau=30} - IV_{\Delta=25,\tau=30}$
PCs	(4) TS hump	$IV_{\Delta=50,\tau=30} + IV_{\Delta=50,\tau=365} - 2IV_{\Delta=50,\tau=91}$
▷ IV portfolios	(5) Smile	$IV_{\Delta=75,\tau=30} + IV_{\Delta=25,\tau=30} - 2IV_{\Delta=50,\tau=30}$
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# PCs and IV portfolios

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# Model Definition and Identification

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## Stock return process

$$\frac{dS_t}{S_t} = R(X_t)dt + Tr \left[ \sqrt{X_t} dZ_t \right] \quad (1)$$

## Matrix-valued factor dynamics ( $X_t \in \mathcal{S}_n^+$ )

$$dX_t = (\Omega\Omega' + MX_t + X_tM')dt + \sqrt{X_t}dW_tQ + Q'(dW_t)'\sqrt{X_t}$$

Correlated innovations:  $Z_t = WR' + B\sqrt{Id - RR'}$

## Pure diffusive case: Wishart Process

Bru (1991):  $\Omega\Omega' = kQ'Q$  (G.T 2008  $k = \text{d.o.f}; k > n - 1$ )

**Focus on  $2 \times 2$  diffusive case,** i.e. a 3-factor model

# Identifiability of Affine models

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We follow the definition of identifiability proposed in (C-D G J JF 2008): a model is identifiable if the state vector and the parameter vector can be identified if the prices for a suitable set of spanning securities are given.

- P1 If  $X_t$  is observable then its drift is observable

$$\mu^{\mathbb{Q}}(t) = \lim_{\Delta \rightarrow 0} \frac{E_t^{\mathbb{Q}} [X_{t+\Delta} - X_t]}{\Delta}$$

- P2 The quadratic variation of an observable process  $X_t$  is observable

$$V_X(t) = \langle X_t, X_t \rangle$$

- P3 The quadratic covariation between two observable processes  $X_t, Y_t$  is observable

$$V_{XY}(t) = \langle X_t, Y_t \rangle$$

# Identifiability in Option Markets

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The state vector is rotated to a set of observable factors. We apply the same identification procedure in option markets. El-Karoui Durrleman(QF 2008) relations identify the relation between 3 fundamental contracts and the infinitesimal characteristics of the factor processes in terms of the implied volatility surface  $IV_t(\tau, K)$ :

$$IV_t(\tau, K) = |\sigma_t| + \tau \mathcal{M}_t + \frac{K - S_t}{S_t} \mathcal{S}_t + O \left\{ \left( \frac{K - S_t}{S_t} \right)^2 \right\}$$

□ The Volatility Level  $|\sigma_t|$ :

$$|\sigma_t| = \lim_{\tau \rightarrow 0} [IV_t(\tau, K)]_{K=S_t} = Tr [X_t]^{1/2}$$

□ The Calendar spread  $\mathcal{M}_t$ :

$$\mathcal{M}_t = \lim_{\tau \rightarrow 0} \left[ \frac{\partial IV_t(\tau, K)}{\partial \tau} \right]_{K=S_t} = \frac{Tr [kQQ^T] - Tr [2MX_t]}{4Tr [X_t]^{1/2}} - |\sigma_t|^2 \mathcal{C}_t$$

# Stochastic Leverage

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- The Risk-Reversal proportional to the skew  $\mathcal{S}_t$ :

$$\mathbf{S}_t = \rho_t \frac{\text{Tr} [QQ^T X_t]^{1/2}}{2\text{Tr} [X_t]} = \frac{1}{2|\sigma_t|^3} \text{Tr} [RQX_t]$$

where  $\rho_t$  the stochastic correlation leverage is given by:

$$\rho_t = \frac{\text{Tr} [RQX_t]}{\text{Tr} [X_t]^{1/2} \text{Tr} [QQ^T X_t]^{1/2}}$$

- The definition of the factors  $|\sigma_t|$  and  $\mathcal{M}_t$  depends solely on the trace, which is a spectral invariant.

**The identification of the matrix coordinate system requires a relation between the dynamical evolution of the reference system and the Risk reversal.**

# Wishart PCA: Eigenvalues

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Consider the case  $n = 2$ ; repeating the computation of Bru (1991) it is easy to compute the evolution of the PCA  $\Sigma_t = U_t \text{diag}(\underline{\lambda}_t) U_t^T$  for the volatility matrix:

$$dS_t = S_t \left( \sqrt{\lambda_1} d\nu_t^1 + \sqrt{\lambda_2} d\nu_t^2 \right)$$

$$d\lambda_i = \left[ k (Q^U Q^{TU})_{ii} - 2M_{ii}^U \lambda_{it} \right] dt + \sqrt{\lambda_{it}} 2 (Q^{TU} Q^U)_{ii} d\nu^i$$

$$- \sum_{j=1}^2 \frac{(\lambda_{it} + \lambda_{jt})}{(\lambda_{jt} - \lambda_{it})^2} \left[ \lambda_{jt} (Q^{UT} Q^U)_{ii} + \lambda_{it} (Q^{UT} Q^U)_{jj} \right] dt$$

$$i = 1, 2$$

# Wishart PCA: Eigenvectors

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The eigenvalues of the Wishart behave as two non-crossing square root processes eigenvectors matrix is a 2-dim orthogonal matrix, hence it is parameterized by a single factor:

$$dA_{12} = \frac{(k (Q_t^U Q_t^{TU})_{12} - M_{12t}^U \lambda_{2t} - \lambda_{1t} M_{21t}^U)}{(\lambda_{2t} - \lambda_{1t})} dt + \frac{\sqrt{(Q_t^U Q_t^{TU})_{12} (\lambda_{1t} + \lambda_{2t})}}{(\lambda_{2t} - \lambda_{1t})} d\beta_t$$

where  $A_t$  is the stochastic logarithm of  $U_t$ :

$$U_t = \begin{bmatrix} \cos(\alpha_t) & -\sin(\alpha_t) \\ \sin(\alpha_t) & \cos(\alpha_t) \end{bmatrix}$$

$dU_t = U_t (dA_t + d\Gamma_t)$ ,  $d\Gamma_t = \frac{1}{2} dA_t dA_t$  and  $d\beta, d\nu_1, d\nu_2$  are independent Brownian Motions  $d\nu_i d\nu_i^S = R_{ii}^U dt$   $i = 1, 2$ .

# Wishart PCA: Summary

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Summarizing, we have a model for returns with

1. Two orthogonal short and long term stochastic volatility factors (two factor Heston models  $\lambda_1, \lambda_2$ ) which cannot cross
2. Factor loadings of the two components evolve stochastically depending on a single factor (the angle  $\alpha_t$ )
3. The factor loadings are independent from the eigenvalues if and only if the  $QQ^T$  is proportional to the identity

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# Estimation

# Parameter Identification

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$\theta = \{M, R, Q, k\} \rightarrow$  13 diffusive parameters

Canonical model  $\rightarrow$  10 parameters to estimate

Some admissibility constraints  $(M, R, k)$

## Two state representations

1. Choleski decomposition:  $X_t = S_t' S_t$

2. Eigenvalue representation:  $X_t = P_t D_t P_t'$

□ Matrix of eigenvalues  $D_t = \begin{pmatrix} \lambda_1 & 0 \\ 0 & \lambda_2 \end{pmatrix}$

□ Eigenvectors in polar coordinates  $P_t = \begin{pmatrix} \cos \alpha_1 & \cos \alpha_2 \\ \sin \alpha_1 & \sin \alpha_2 \end{pmatrix}$

$\alpha_2 = \alpha_1 + \pi/2 \rightarrow$  symmetry, rotation matrix

**Two representations:**  $\{s_{11}, s_{12}, s_{22}\}$  or  $\{\lambda_1, \lambda_2, \alpha\}$

# Maximum Likelihood estimation

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▷ ML estimation

ML estimation 2

ML estimation 3

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Joint estimation procedure, slightly simplified version of Bates(2000)

- Parameters  $\theta = \{M, R, Q, k\}$
- State  $\{X_t\}$ ,  $t = 1, \dots, 59$
- Unconditional error covariance matrix  $Q$  to calculate heteroskedasticity correction  $\Omega_t$ ,  $t = 1, \dots, 59$  for

$$\max_{\theta, \{X_t\}} \ln L = -\frac{1}{2} \sum_t \ln |\Omega_t| + e_t' \Omega_t^{-1} e_t$$

Note: dimension of  $\Omega_t$  changes daily, depending on number of observations.

# Maximum Likelihood estimation 2

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ML estimation

▷ ML estimation 2

ML estimation 3

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**Step 1:** Nested NLS estimation to get starting values for ML

$$\hat{\theta} = \arg \min_{\theta} \left( \hat{C}_i(\theta, X_t^*(\theta)) - C_i \right)^2 \quad (7)$$

$$X_t^*(\theta) = \arg \min_{X_t} \left( \hat{C}_i(\theta, X_t) - C_i \right)^2 \quad (8)$$

**Step 2:** Heteroskedasticity correction

Bin data in 4 moneyness and 3 maturity bins.

Calc. unconditional within/between bins error cov. matrix  $Q$ .

Produce daily unconditional error covariance matrix  $\Omega_t$

**Step 3:** Perform ML parameter estimation.

State still estimated via NLS.

Iterate between step 2 and step 3.

# ML estimation 3

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**Step 4:** Error estimation for  $\theta$  using BHHH algorithm  
Gradients calculated numerically.

$$\hat{V}_{BHHH} = \left( \frac{1}{N} \sum_i s_i(\hat{\theta}) s_i(\hat{\theta})' \right)^{-1} \quad (9)$$

# Out of sample

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## Semi-out of sample extension

Add daily data within the sample period (59 → 1254 obs)

Use parameters  $\hat{\theta}$  from monthly data

Estimate time series of daily state via NLS.

Virtually same performance.

## Full of sample extension

Furthermore add daily data from 1996-1999 and 2005-mid2006.

(→ 2850 obs)

Notably worse performance.

<i>rms</i> \$ errors	monthly data	semi-OOS	ful-OOS
Full Wishart	1.059	1.099	1.547

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Parameter est.

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# Parameter estimates + interpretation

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$$m_{11} \quad -0.2906 \quad (0.0218)$$

$$m_{21} \quad -9.3628 \quad (0.0603)$$

$$m_{22} \quad -5.0719 \quad (0.0180)$$

- Distinctly different mean reversion speeds.
- Long-run factor  $\lambda_1$ , ca. 3yrs
- Short-run factor  $\lambda_2$ , ca. 2 months
- Very fast mean reversion of  $U_t$

$$q_{11} \quad 0.0340 \quad (0.0049)$$

$$q_{22} \quad 0.4789 \quad (0.0079)$$

- Volatility of short-term shocks is larger.

# Parameter estimates (2)

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$$R_{11} \quad -0.9757 \quad (0.1362)$$

$$R_{12} \quad 0.1586 \quad (0.0335)$$

$$R_{22} \quad -0.6905 \quad (0.0186)$$

- Different contributions of short- and long-run skew.
- Positive  $R_{12}$  allows for change of sign in correlation leverage

$$\beta \quad 1.0000 \quad (0.0427)$$

Data: SP500 call options, 59 trading days in 2000-2004 (2nd Wednesday before expiry), overall 9029 observations

# Results: State

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Parameter est.

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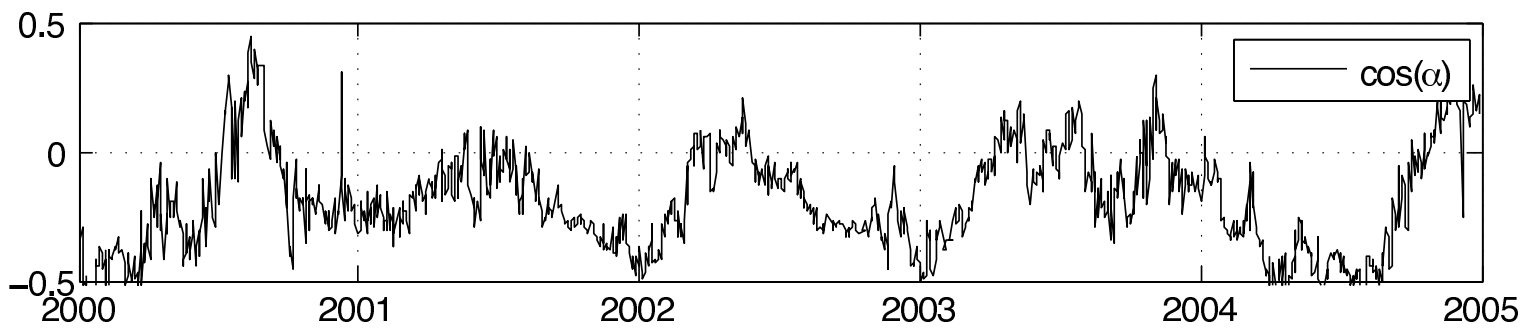
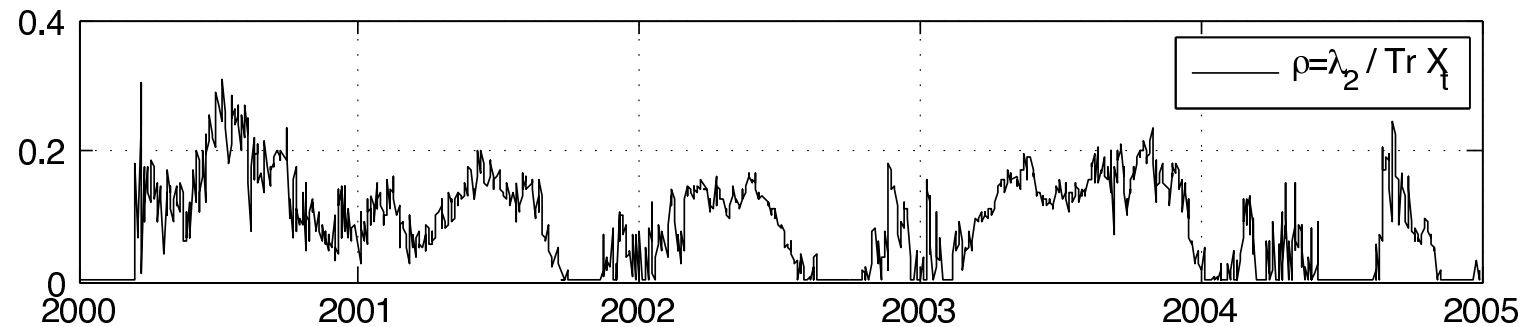
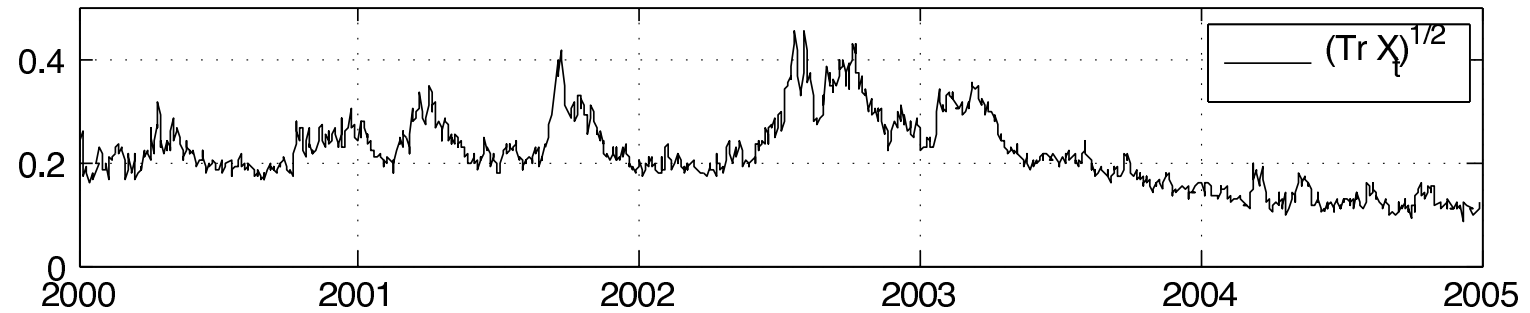
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- Short run factor  $\ll$  long run factor
- $U_t$  changes sign

Next goal: find an interpretation of the state in our framework of

- Level
- Term structure
- Risk reversal

# Level – no magic

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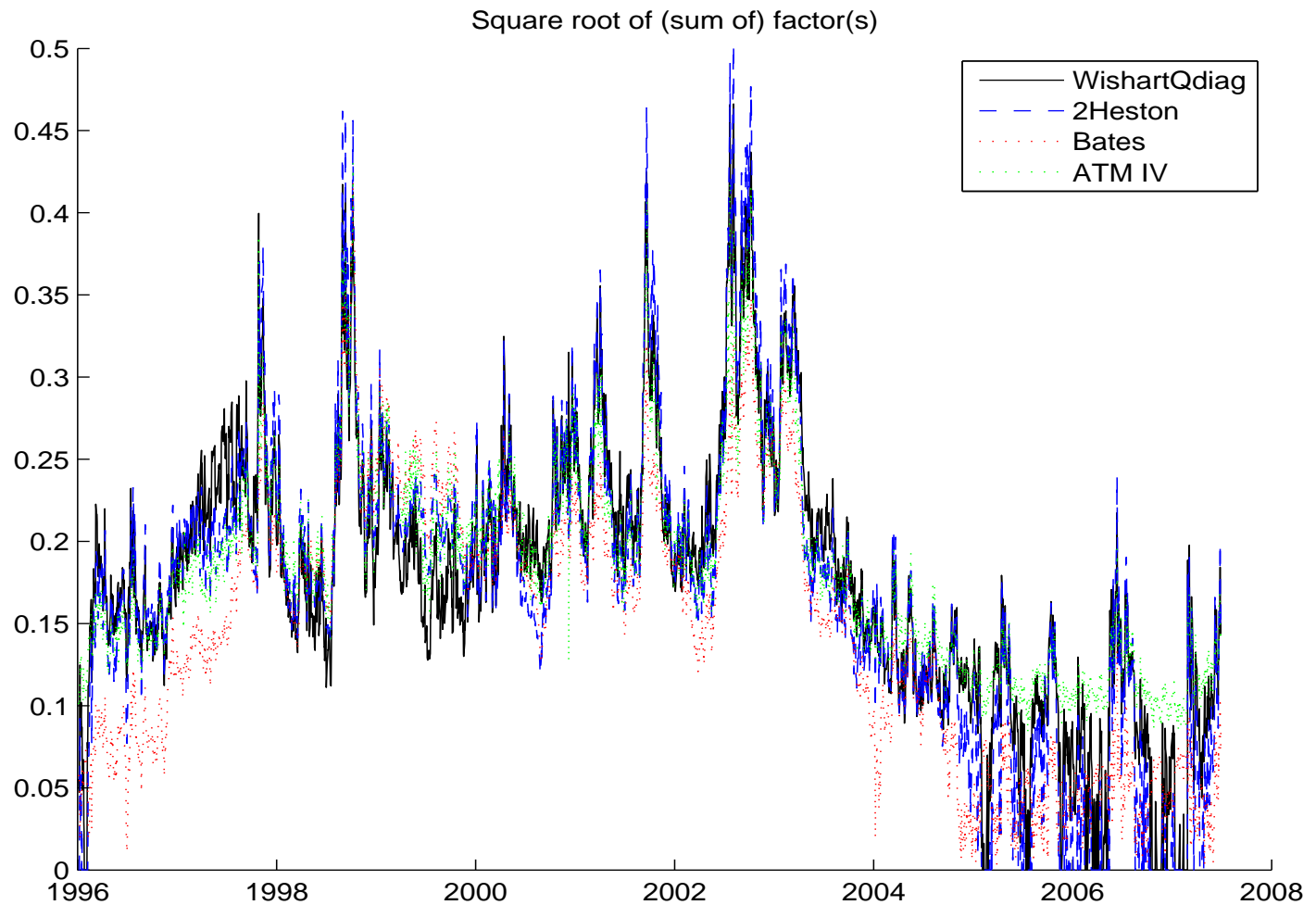
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# Term structure

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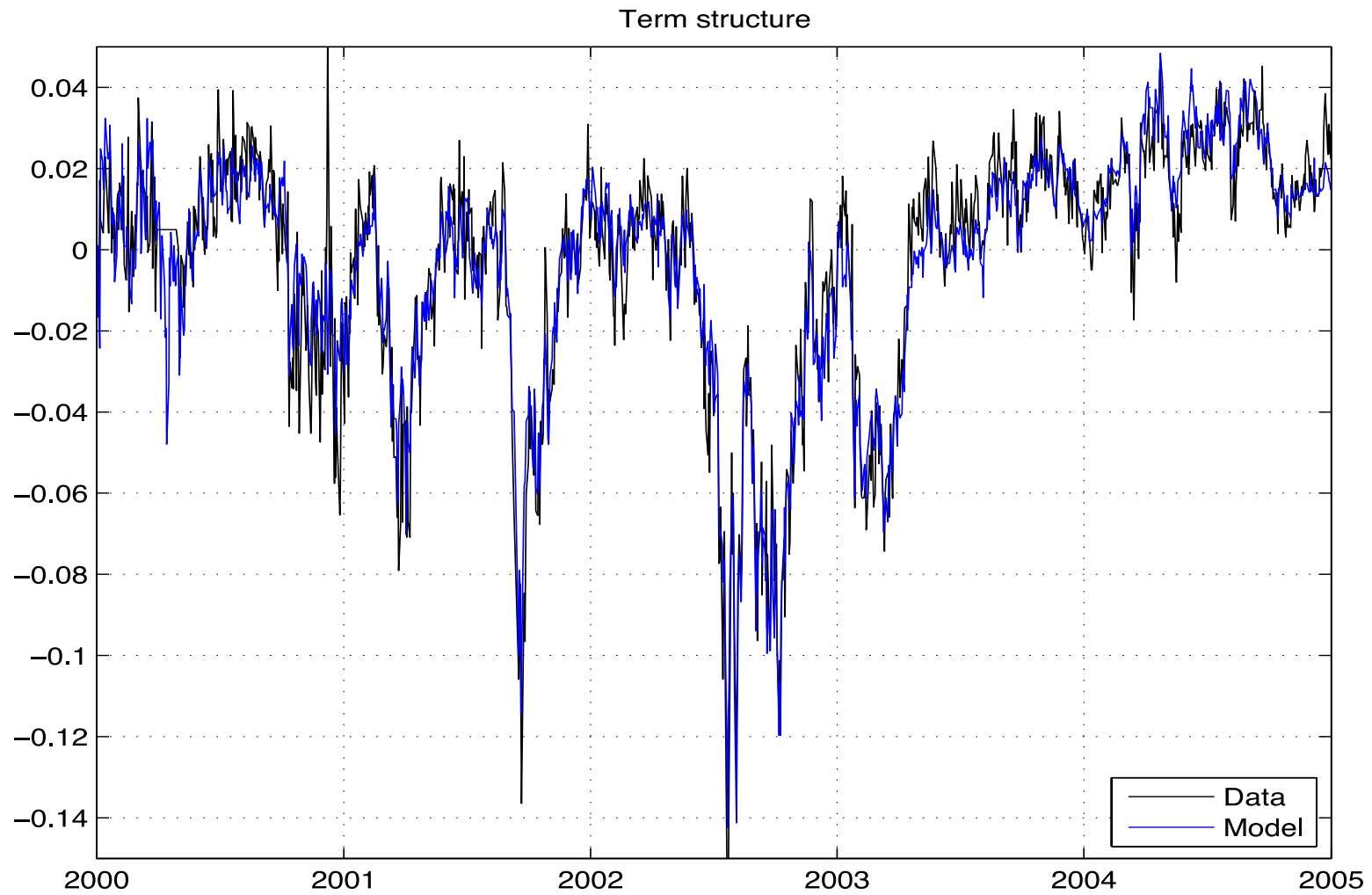
Conclusion

$$\text{Term structure} \approx \alpha + \beta_1 V_t + \beta_2 \lambda_t^1 / V_t$$

Heston-like effect links term structure to level

Mixing of different mean-reversion speeds for  $\lambda_t^1$  and  $\lambda_t^2$  allows for deviation

# Term structure



$$R^2 = 0.81$$

# Risk Reversal

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- Interpretation
- State (eigenval)
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A stylized model from estimation results:

$$Q \cdot Q^T = \begin{bmatrix} 3.7265 \times 10^{-4} & -5.6389 \times 10^{-3} \\ -5.6389 \times 10^{-3} & 0.16941 \end{bmatrix} \simeq \begin{bmatrix} 0 & 0 \\ 0 & 0.16941 \end{bmatrix}$$

$$RQ = \begin{bmatrix} -6.6341 \times 10^{-3} & -0.23616 \\ 0 & -0.25816 \end{bmatrix} \simeq \begin{bmatrix} 0 & 1 \\ 0 & 1 \end{bmatrix} (-0.24716)$$

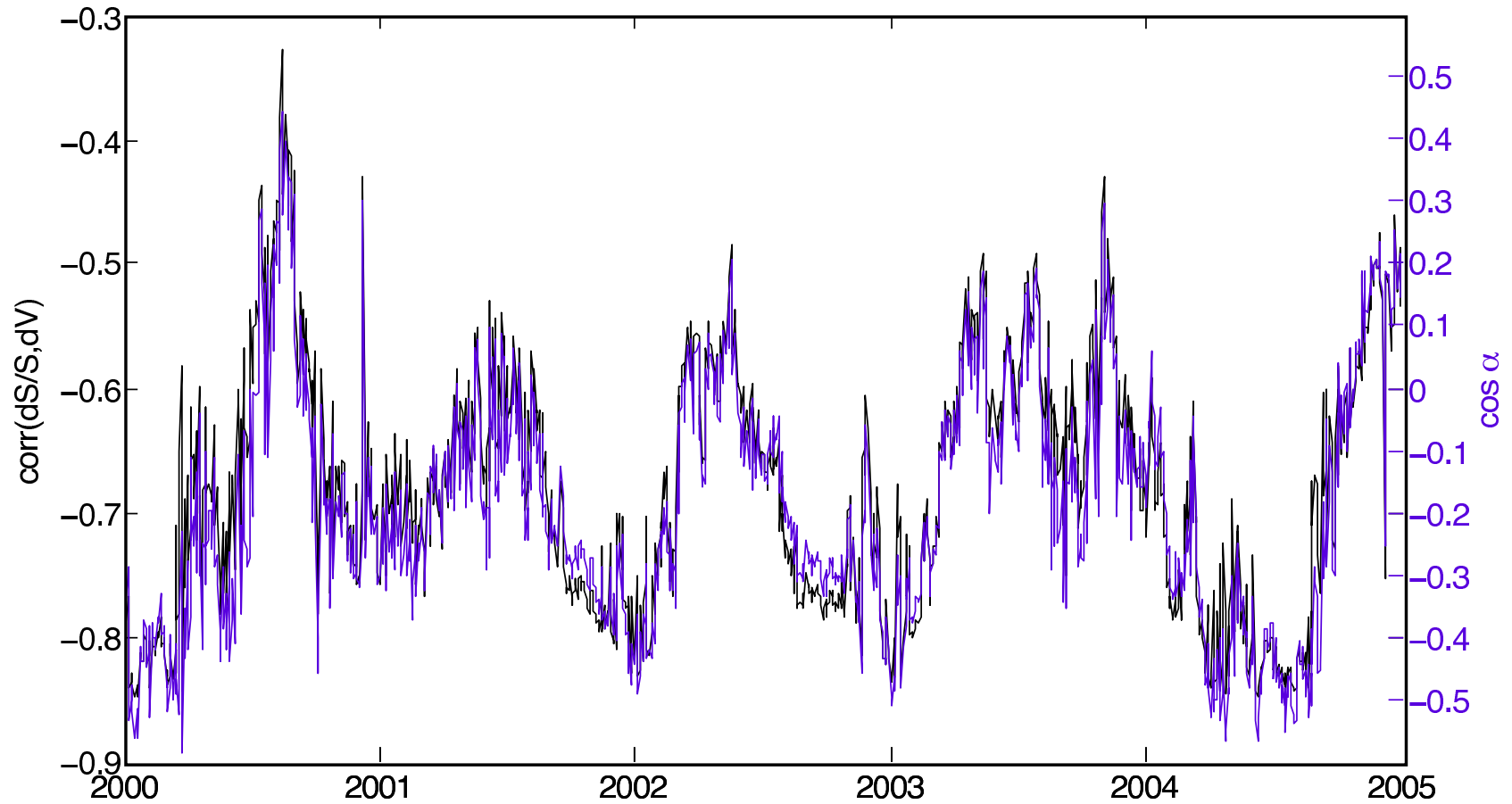
$$\begin{bmatrix} \sqrt{\lambda_1^\infty} & 0 \\ 0 & \sqrt{\lambda_2^\infty} \end{bmatrix} = \begin{bmatrix} 2.5551 \times 10^{-2} & 0.0 \\ 0.0 & 0.14945 \end{bmatrix}$$

$$\frac{Tr [RQ \Sigma_t]}{\sqrt{Tr [\Sigma_t]} \sqrt{Tr [Q^T Q \Sigma_t]}} \simeq -0.60920 [1 + \cos \alpha_t]$$

**The rotation factor is directly proportional to a stochastic skewness component orthogonal to the volatility factors!**

# Stochastic leverage – $\cos \alpha$

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Related literature

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# Performance comparison

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## Data:

- S&P 500, monthly data
- Sample period 2000-2004 (59 observations)
- Time to maturity:  $10d \leq \tau \leq 1yr$
- Minimal cuts in moneyness:  $\min(C_i, P_i) > \$0.375$

Model	Factors	<i>rms</i> \$ error	remark
Heston	1	1.676	
Bates (1996)	1	1.534	State-ind. jumps
Christoffersen (2008)	2	1.202	Feller cond. violated
Full Wishart	3	1.059	

# Conclusion+Outlook

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Related literature

- Identifiability of the Wishart Multifactor Volatility is obtained in terms of observable portfolios
- Good cross-sectional and time-series performance with one parameter set for 5 years of data
- Work in progress: computation of estimation errors, addition of Jumps

# Related literature

## Empirical facts on the IV surface

Skiadopoulos, Hodges, Clewlow (2000): The dynamics of the S&P 500 implied volatility surface.

Fengler, Härdle, Villa (2001): The dynamics of implied volatilities: A common principle components approach.

## Affine processes and Transform Analysis

Carr and Madan (1999): Option Valuation Using the Fast Fourier Transform

Duffie, Pan, Singleton (2000): Transform analysis and asset pricing for affine jump-diffusions

Grasselli and T. (2008): Solvable Affine Term Structure models

Leippold, Trojani (2008): Asset Pricing with Matrix Affine Jump Diffusions

## Wishart process + applications

Gourieroux, Suffana (2004): The Wishart Autoregressive Process of Multivariate Risk

Buraschi, Porchia, Trojani (2008): Correlation Risk and Optimal Portfolio Choice

## Multifactor option pricing models

da Fonseca et al.(2008): A Multifactor Volatility Heston Model

Christoffersen et al.(2008): The Shape and Term Structure of the Index Option Smirk: Why Multifactor Stochastic Volatility Models Work so Well

Christoffersen et al.(2008): Option Valuation with Long-run and Short-run Volatility